

Research Article

Effect of Implementing a Free Delivery Service Policy on Women's Utilization of Facility-Based Delivery in Central Ethiopia: An Interrupted Time Series Analysis

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Background. Access to and utilization of facility delivery services is low in Ethiopia. The government of Ethiopia introduced a free delivery service policy in all public health facilities in 2013 to encourage mothers to deliver in health facilities. Examining the effect of this intervention on the utilization of delivery services is very important. **Objective.** In this study, we assessed the effect of provisions of free maternity care services on facility-based delivery service utilization in central Ethiopia. **Methods.** Data on 108 time points were collected on facility-based delivery service utilization (72 pre- and 36 postintervention) for a period of nine years from July 2007 to June 2016. Routine monthly data were extracted from the District Health Information System and verified using data from the delivery ward logbooks across the study facilities. An interrupted time-series analysis was conducted to assess the effect of the intervention. **Results.** The implementation of the free delivery services policy has significantly increased facility deliveries. During the study period, there was a statistically significant increase in the number of facility-based deliveries after the 24th and 36th months of intervention ($p < 0.05$). Program effects on the use of public facilities for deliveries were persisted over a longer exposure period. **Conclusion.** The findings suggested that the provision of free delivery services at public health facilities increased facility delivery use. The improved utilization of facility delivery services was more marked over a longer exposure period. Policy-makers may consider mobilizing the communities aware of the program at its instigation.

1. Introduction

Improving maternal service utilization to attain sustainable development goals and to reduce the noticeably high maternal and neonatal mortality in low-income countries remains a pressing priority [1]. Maternal mortality is generally high in Sub-Saharan African (SSA) countries accounting for 66% of maternal deaths globally [1]. Ethiopia is one of the countries with high maternal mortality, 412 per 100,000 live births [2]. Efforts to end preventable maternal and neonatal deaths consider universal coverage of facility delivery services as one of the key interventions [1, 3].

Facility-based delivery is vital in the reduction of maternal mortality; attaining high coverage alone can prevent up to 54% of maternal deaths [4]. Yet only 46% of women in SSA gave birth in health facilities [5]. Thus, to increase the utilization of facility delivery, many African countries have initiated free maternity services (FMS) [6, 7]. Studies across many SSA countries generally suggest increased utilization of facility delivery with the removal of user fees [5, 8, 9].

A study by McKinnon et al. for a number of SSA countries confirms that the removal of user fees is associated with an increase of 3.1 facility-based deliveries per 100 live births [7].

In Zambia, following the removal of user fees, facility-based deliveries did not increase significantly for public facilities, but a significant positive effect for private (faith-based) facilities with limitations because it only accounts for five percent of the sample [10]. A study in Ghana showed variation in coverage of delivery fees between districts [11].

In Kenya, recent evidences related to facility-based deliveries are heterogeneous. One article mentions a significant increase in the number of deliveries over a longer exposure period after abolition [12], while another reports that the increase in the number of deliveries was observed during the early phase of the program [13]. The instance of unsustainable increase in facility delivery service utilization following the removal of user fees was also reported in Nepal's study [14, 15].

However, only a few studies have used rigorous analysis adjusting for underlying trends so far [6, 7, 16, 17]. Hence, evidence on the effect of free delivery service policy on utilization facility delivery remains limited and mixed.

In Ethiopia, despite the implementation of the free delivery services policy in the country since 2013 [18], to date, 26% of mothers delivered in health facilities [2]. Since the initiation of the free maternal care services in Ethiopia, no study has been undertaken to assess its effect on the utilization of health facilities for delivery services. Thus, this study was conducted to assess the effect of the provisions of FMS on the utilization of facility-based delivery services in the East-Shewa Zone, Central Ethiopia.

2. Methods

2.1. Study Setting. The study was conducted in the East-Shewa zone, one of the sixty-eight zones in Ethiopia. The Zone was selected because it was one of the administrative zones that began implementing the national free delivery service intervention at the initiation phase in 2013 in all public health centres [18]. The Zone had an estimated population of 1,967,077 inhabitants based on the projected population of Ethiopia, 2017 [19].

The Ethiopian health care system encourages normal deliveries to be conducted at health centers, and the need to refer only complicated cases to hospitals [18].

The government has created a primary level care consisting of health posts (HPs), health centers (HCs), and primary/district hospitals which are all connected to each other through a referral system [18]. A free ambulance service is organized in every district to facilitate referral for obstetrics emergencies [20].

2.2. Study Design. The study utilized an interrupted time series design, which is appropriate to compare changes in the utilization of facility delivery services before and after the introduction of the free service intervention [21–23]. Given that there was a definite point in time when the implementation of the policy began and data were available in the health facilities for the time period before and after the intervention, interrupted time-series design is the ideal method.

2.3. Study Facilities and Sampling Strategy. The process of selecting health facilities had two phases: we first considered

all of the 25 HCs within the zone. Then, we made a preliminary assessment and initially identified 13 HCs within the zone that had complete facility birth data 2007–2016; the rest 11 had fragmentary delivery ward logbooks in place, and the data were confused with omitted months, year, torn page, and multiple outliers and one HC (Methehara) had only 3 years data 2014–2016 because of flood in 2013. Hence, the 12 HCs were excluded from the sampling frame.

In the second phase, of the 13 HCs which had the last nine years' complete data, 5 HCs were randomly selected. A flow diagram showing the steps taken for selecting facilities is presented in Figure 1.

2.4. Data Collection. The outcome variable was the total number of deliveries taking place in public health facilities per month. Information on the intervention period (both before and after the launch of the policy on July 1, 2013) was obtained from the East-Shewa health office.

Data were extracted from district health records using a structured data retrieving form. These were then verified using facility-based delivery ward logbooks to ensure data accuracy in the study. The data were collected for the period between July 2007 and June 2016, a total of 108 months, 72 pre- and 36 postintervention months. Experienced record officers were recruited, trained, and used in the data collection. The first officer collected the data from each study site. The second officer counter checked and verified whether the collected data were similar to what was in the study facilities' logbooks. Only verified data were used for analysis.

2.5. Data Analysis. The data were verified and cleaned using an Excel spreadsheet then exported to Stata version 12 for analysis. We carried out an interrupted time series (ITS) analysis (segmented regression model) to estimate the effect of the policy on facility deliveries [17] (clean data found as supplementary file 1). The segmented regression model used for the study gives three coefficient β_1 , β_2 , and β_3 :

$$Y_t = \beta_0 + \beta_1 T_t + \beta_2 X_t + \beta_3 X_t T_t + \varepsilon_t, \quad (1)$$

where Y_t is the number of facility-based deliveries (FBD), β_0 estimates the baseline level FBD at the beginning of the period, β_1 estimates the change in the number of FBD until the free delivery service policy, β_2 is the change in FBD immediately after the free delivery service policy, and β_3 is the difference in the trends of FBD between pre- and post-changes in policy (effect of the intervention over time).

The covariates are defined as follows: T_t is the time from the start to the end of the observation period and X_t is a dummy variable coded 0 and 1 for periods preintervention (from July 2007 to June 2013) and postintervention (July 2013 to June 2016), respectively; $X_t T_t$ is an interaction term between the time and intervention dummy; and ε_t is the error term.

Estimates for regression coefficients corresponding to two standardized effect sizes are obtained: a change in level (step change, i.e., β_2) and a change in trend before the intervention (β_1). The change in trend after the intervention (β_3) is the sum of the preintervention slope and the change in

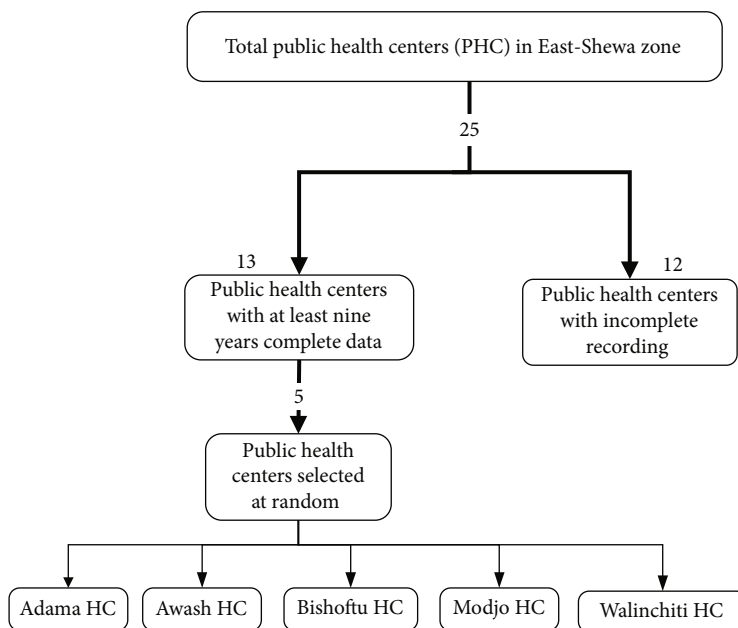


FIGURE 1: A flow diagram showing the steps taken for selecting the facilities within the East-Shewa zone.

level. Slope and level regression coefficients were important in calculating the 95% confidence intervals. The p values demonstrated the significance of the effect of the free delivery service policy. Percentages were calculated to estimate the relative effect of the policy.

Policy effect equates $= \text{FBD with policy} - \text{FBD without policy}$, where $\text{FBD without policy} = \beta_0 (\text{constant}) + \beta_1 T_t$ and $\text{FBD with policy} = \beta_2 \times X_t + \beta_3 \times X_t T_t$.

We estimated the absolute effect of the free facility-based delivery policy at 12, 24, and 36 months or at the 1st, 2nd, and 3rd year of implementation, respectively, with the derivative of the policy shown here $\text{FBD} = \beta_2 \times X_t + \beta_3 \times X_t T_t$.

The main aim of the regression analysis here is to measure the outcome of interest by keeping out the influence of potential time-varying confounders on the effect of the free delivery service intervention [22–24].

The analysis was done to determine the twelfth-month level effect after the policy execution date to monitor any short-term effects of the policy. The analysis for the twenty-fourth month was to demonstrate any midterm effects, whereas the analyses done at the thirty-sixth-month postintervention were to assess whether the policy had a long-term effect. Statistical significance was set at $p \leq 0.05$.

3. Results

Using the data at the district level, we compiled zonal data for the 5 districts for a period of 9 years, 2007–2016. The list of omitted facilities is presented as supplementary file 1.

The population of women of reproductive age in the districts pre-post ranged from 7,119 vs. 8,696 at Modjo district to nearly eight times that in Adama, and the expected pregnancies ranged from 330 to 2,556 pregnancies per year across districts in the zone.

Table 1 depicts the population of women of reproductive age and service data before and after the free delivery care policy implementation. There was an increase in facility deliveries within the zone and across facilities except for the Adama health center.

Figure 2 shows the total number of monthly facility deliveries in the study health facilities. There was an immediate increase in the number of facility-based deliveries just following free delivery services policy implementation and then a temporary drop. The observed effect was smoothed by ITS analysis that involves a before-after comparison within a single population by creating control variables to estimate the counterfactual results, rather than a comparison with a control group. A trend was established for the before and after implementation of the intervention for facility-based deliveries with a goal of determining whether the intervention had a significantly greater effect than any underlying secular or seasonal trends. Throughout the postintervention period, the observed number of FBD under the free delivery services policy introduction was consistently higher than the expected number under the counterfactual results of nonintroduction of the free delivery services.

Table 2 shows the level and trend effects and the relative effects of the policy by year after the intervention. There had been a gradual increase of 0.8 facility-based deliveries per month before the implementation of free service (95% CI: 0.213–1.4; $p < 0.01$). Following the implementation of the free delivery services policy, there was an immediate increase of 320 facility-based deliveries conducted within health institutions ($p < 0.05$); 44% of this increase is directly attributable to the policy. There was a long-term effect of the policy by 350 (51% of which was due to the policy) more deliveries occurred in the third year ($p < 0.001$).

Thus, the average number of facility deliveries increased to 77, 98, and 118 per month in all facilities; in other words,

TABLE 1: The population of women of reproductive age and service data before and after the free delivery care policy implementation.

	Preintervention	Postintervention
Population data	July 2007-June 2013	July 2013-June 2016
Zonal women of reproductive age	2,169,190	1,261,067
Expected deliveries	82,429	47,920
Average expected deliveries per month	1,145	1,331
Total facility-based deliveries (FBD) (ratio)	14,270 (17.3%)	10,460 (21.83%)
Average facility deliveries per month	198	291
FBD by health facilities		
Adama health center	6,686	3,138
Awashmelkasa health center	850	1,905
Bishoftu health center	863	1,488
Modjo health center	3,991	1,761
Walinchiti health center	1,880	2,167

36, 44, and 51 of any 100 births in the health facilities could be attributed to the free delivery service implementation, compared with what it would have been without the free delivery service. These showed that the average number of facility deliveries increased by 36%, 44%, and 51% following the intervention at the 1st, 2nd, and 3rd year of implementation, respectively. However, the increment in the 1st year of the implementation was not statistically significant (Table 2).

4. Discussion

The implementation of the free delivery service policy in public health facilities in the study zone in 2013 was associated with a statistically significant increase in the utilization of facility-based deliveries. This finding is consistent with the results of the implementation of free maternal health care policies across several African countries [5, 7–9, 16]. A statistically significant increase in the use of public facilities for deliveries remained consistently high over the 2 years of exposure to the program. The effect size was also higher in longer phases. The high level of utilization of free delivery services over a long period of time in this study creates an opportunity to address the observed maternal and neonatal death toll.

This finding is in contrast with other studies, in which increased utilization of free delivery care services was documented during the early phase of the program implementation [6, 18].

Unlike the longer phase, the result for the likelihood of delivery in a public facility was statistically insignificant in the early phase, even though positive in sign in this study.

One of the reasons for observation of an insignificant effect in the early phase of the program might be low levels of awareness in the larger population about the availability of free delivery services. For example, in Ghana, communities apparently did not properly understand the abolition measure, despite various types of publicity effort [25]. Further, a qualitative finding conducted in three districts of Kenya supports these ideas, noting multiple challenges in program implementation, including inadequate stakeholder engage-

ment and confusion on eligibility criteria and lack of reimbursement to health facilities for providing free delivery care.

It is possible that the free delivery services policy, despite being announced, was not completely internalized by either the health professionals or expectant mothers. Low levels of awareness in the larger population in the early phase may also explain the findings over the longer exposure period. The findings suggest the need to mobilizing communities (occasioned by women's group called "Women's Health Development Army (WHDA)") [26–28] to be aware of the availability of free delivery and a 24-hour ambulance services for laboring women [29].

There have been improvements in the utilization of maternal health services as expected; studies have shown that facility deliveries usually increase following similar fee exemption policies [10, 11, 15, 16].

Although the rise in facility-based deliveries within the study zone is promising, considerable work remains to make the communities initially aware of the availability of free maternal services and to demand utilization. Studies suggest many possible explanations that could range from low coverage of the free delivery services due to overflow of users vs. lack/delays of reimbursement or that the policy might not evenly be implemented across treated districts [9, 30, 31].

Numerous studies showed that user charges are not the only financial determinants of seeking delivery services—unavailability and high costs of transport services and other delivery-related service costs are also financial barriers to access delivery services [32–34].

Further, variations in the attribute of free delivery service results are expected, since user fees are not the only factors to determine the use of facilities for delivery, and therefore, the relative importance of the fees compared to other determinants is likely to vary from time to time. Thus, the potential impact and magnitude of the policy change are likely to also vary accordingly [9].

The strengths of the current study include that data were extracted over a long period of time to absorb the effect of yearly fluctuations, the quality of data obtained from district records was verified by reviewing the delivery ward logbooks

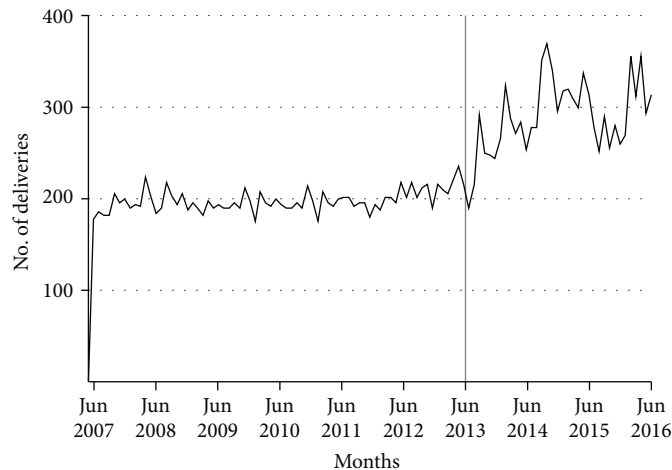


FIGURE 2: Trends for facility-based deliveries before and after implementation of the free maternity services policy in East Shewa zone (the line at 72 months).

TABLE 2: Parameter estimates, standard errors, and p values of a segmented regression model predicting the average number of deliveries per month in East-Shewa zone, central Ethiopia, 2007-2016.

Independent variables	Coefficient	Standard error	p value	RE%
Model (correcting for the first-order autocorrelation)				
Initial level of facility deliveries (β_0)	173.867	8.791	0.000***	
Preintervention slope (β_1)	0.807	0.299	0.008**	
Immediate intervention effect (β_2)	56.793	10.934	0.000***	
Postintervention slope (β_3)	1.698	16.817	0.009**	
Level effect				
1 st year	290	16.542	0.452	36.2
2 nd year	320	11.828	0.015*	44
3 rd year	350	10.934	0.000***	51

Note: *** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$.

by trained and expert data collectors, and the assumptions for interrupted time series analysis were fulfilled. These are clearly differentiated time periods and sequential measures with a clear outcome of interest [22].

While the use of an interrupted time series (ITS) design is a strength of the study, it also presents a limitation, in that it is not possible to identify contextual factors associated with differences in the magnitude of service utilization.

The free maternal health care policy has been implemented in all public health facilities. This brought limitations in getting control groups (public health facilities where the policy was not implemented) for comparisons. Although most of the findings from this study are consistent with other studies, externalization of findings to depict the national picture in the implementation of the free maternal health care policy may be questioned given the shortcomings to address other contextual factors affecting health facility delivery services utilization in this study.

Further, it is possible that the data abstracted manually by the record officer are susceptible to reporting bias. This risk may be smoothed by the fact that data were verified using health facility records.

The fact that this study demonstrated free maternity services can significantly increase the number of facility delivery shows cost is one of the major barriers to facility-based delivery service utilization; however, it may not be the only factor contributing to the utilization of health facility for delivery as the majority of mothers in the study areas still deliver at home. Therefore, there is a need to consider other determinants that might be very crucial for the creation of delivery service demand by mothers. To comprehensively understand the effect of the free delivery policy, future qualitative studies should be carried out to understand other determinants in the study area with the context of the free delivery care policy.

Consistency in the direction of effects suggests robust evidence that user fees have an effect on the utilization of health facility delivery services in this study.

5. Conclusion

The findings suggest that the provision of free delivery services at public health facilities increased facility delivery use. The improved utilization of facility delivery services

was more marked over a longer exposure period. The high level of utilization of free delivery services over a long period of time in this study creates an opportunity to address the observed maternal and neonatal death toll.

This finding implies that removal of user fees for facility delivery service utilization over a long period of exposure may serve as an important public health measure to address the observed low level of facility delivery and thereby reduce the highest maternal and neonatal death toll.

Thus, policy-makers may consider mobilizing communities to be aware of the program at its instigation.

Abbreviations

ANC:	Antenatal care
BEOC:	Basic emergency obstetrics care
HC:	Health center
HPs:	Health posts
MCH:	Maternal and child health
MDG:	Millennium development goals
PHCU:	Primary health care unit
SSA:	Sub-Saharan Africa
SDGs:	Sustainable development goals
WHDA:	Women's health development army.

Data Availability

The data supporting the conclusions drawn in this study have been included in the article. However, the datasets underlying the findings of our study are contained within the supplementary materials, as supplementary file 2.

Ethical Approval

This study was approved by Arsi University's Health Research Ethical Review Committee, accreditation number ERC 0122/2012. There was no direct human contact with the respondents in this study. Participant reports/information was already anonymized and de-identified prior to data collection.

Consent

Since this study involved data abstraction, no written consent was obtained from the study participants for using their records.

Conflicts of Interest

The authors have declared that no competing interests exist.

Authors' Contributions

AD, YB, and AW conceived and designed the study, performed data analysis, and wrote the paper. AD was responsible for the data collection. All authors read and approved the final manuscript.

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Supplementary Materials

Supplementary file 1: Table showing the list of health centers that removed user fees in July 2013 and those marked "X" were incomplete and discarded. Supplementary file 2: Table showing a clean dataset used for the analysis of facility delivery in East-Shewa zone, 2007-2016. (*Supplementary Materials*)

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