

## Research Article

# A Research on Modification Method for NSM FRP-Concrete Bonded Joints Strength Models

Jing Li,<sup>1</sup> Liwen Zhang ,<sup>1</sup> Yanggui Deng,<sup>2</sup> and Junping Zhang<sup>1</sup>

<sup>1</sup>Department of Civil Engineering, Guangzhou University, Guangzhou, China

<sup>2</sup>Zhongdu Engineering Design Co., Ltd., Guangzhou, China

Correspondence should be addressed to Liwen Zhang; lwzhang@gzhu.edu.cn

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A modification method was proposed for near-surface mounted (NSM) fiber-reinforced polymer (FRP)-concrete bonded joints strength prediction models considering model uncertainty. A database consisting of 246 test records was involved. Three bonded joints strength prediction models for NSM FRP reinforcement system were selected for modification. All the three selected models have model uncertainty factors associated with input design parameters. Spearman correlation analysis was used to prove the systematic correlation of the model uncertainty factors. For each model, a regression function  $f$  was established to eliminate the systematic nonrandom part of the model uncertainty factor. Then, the model uncertainty factors could be described by random variables obeying logarithmic normal distribution. A reliability analysis using the JC method was carried out to validate the practical significance and value of model modification. This study improves the predictability of FRP NSM reinforcement systems and provides valuable references for model calibration in practical engineering.

## 1. Introduction

As one of the most effective techniques for the strengthening of aged concrete structures [1, 2], near-surface mounted (NSM) reinforcement involves inserting a reinforcing material into the concrete cover of the structural member which needs to be reinforced. In the past decade, researchers have been working on various studies of the NSM FRP reinforcement techniques, promoting its widespread application [3, 4]. For FRP NSM reinforcement systems, the utilization of FRP's mechanical properties and reinforcement effects depends mainly upon their bonded joints strength [5–7].

Some prediction models have been proposed for NSM FRP-concrete bonded joints strength [2, 8–11]. Model uncertainty inevitably exists due to limited experimental data, incomplete research parameters, and idealized calculation methods, resulting in a certain error between the predicted value and the experimental value. From the perspective of engineering application, the calculated strength less than the measured value is regarded to be

conservative, while the opposite is unsafe [12]. At present, the study on how to calibrate existing NSM FRP-concrete bonded joints strength models is still insufficient, unsystematic, and superficial, requiring relative further research [13–16]. Thus, our study is aimed at proposing a calibrating method for commonly used NSM FRP-concrete bonded joints strength models with the consideration of model uncertainty.

A model uncertainty factor was defined and adopted to quantitatively describe the model uncertainty [17]. The key is that this model uncertainty factor must be “random” and should have no dependence on the input design parameters [18]. For some NSM FRP-concrete bonded joints strength prediction models, the model uncertainty factor has a dependency on the design parameters, which does not match the definition. Hence, a regression analysis is needed to eliminate the systematic part from the model uncertainty factor, and the randomness of regression residual needs to be verified.

In this paper, a total of 246 effective NSM FRP pullout test data were gathered. Three NSM FRP-concrete bonded

joints strength models presented by Seracino et al. [19–21], which were confirmed to have model uncertainties, were incorporated for model calibration. The systematic effect on the model uncertainty factor was eliminated by carrying out a multiple regression analysis. The residual part was checked for randomness, and three modified models were obtained. This research can improve the prediction precision of NSM FRP reinforcement systems and provide valuable references for model calibration in practical engineering.

## 2. Bonded Joints Strength Models

A number of prediction models have been applied to calculate the NSM FRP-concrete bonded joints strength. Three commonly used NSM FRP-concrete bond stress models with model uncertainty were incorporated in this paper. They were proposed by Seracino et al. [16], Ali et al. [20], and Zhang et al. [21] as listed in Table 1 and were abbreviated as SR Model, AM Model, and Zhang Model hereinafter. In this paper, the bonded joints strength was defined as the maximum load in a pullout test [22]. Hence, for the three selected bonded joints stress models, peak loads (bonded joints strength) were calculated according to the bonded joints strength model presented by Seracino et al. [19], as shown in the fifth row in Table 1. In this table,  $\tau_{\max}$  is the peak bond stress,  $f_c$  is the concrete compression strength,  $b_p$  is the width of groove,  $t_p$  is the height of FRP,  $\gamma$  is the height-to-width ratio of groove,  $\varphi_f$  is the length-to-width ratio of failure surface,  $P_{IC}$  is the peak bond load,  $\delta_{\max}$  is the maximum bond slip,  $L_{\text{per}}$  is the perimeter of failure surface, and  $(EA)_P$  is the stiffness of FRP.

## 3. Experimental Data

Research shows that the bonded joints strength of a FRP NSM reinforcement system is concerned with a variety of factors, e.g., physical dimension and material property [23]. However, it is impractical to take all relevant factors into account [24]. In this paper, five significant influence parameters were selected as the key factors: (1) concrete compression strength  $f_c$ , (2) FRP modulus of elasticity  $E_f$ , (3) FRP height  $t_p$ , (4) FRP thickness  $t_f$  and (5) groove width  $t_g$  [25].

A database covering all the five selected key factors and the related bonded joints strength was collected for analysis [26, 27]. Five typical failure modes, occurring on three intermediate materials (FRP, adhesive, and concrete) as well as two interfaces (adhesive-concrete and adhesive-FRP), were incorporated into this paper. According to the critical region where failure may occur, they were, respectively, failures in the interface of concrete and adhesive, in the interface of adhesive and FRP, within the adhesive, in a single material (concrete crushing or FRP rupture), and in the surface of concrete (cracks propagate through the concrete). Of particular note is that this study only took pullout tests into account instead of bending tests here [28]. Because the bonded joints strength of a FRP NSM reinforcement system is normally evaluated by a pullout test.

TABLE 1: Selected bonded joints strength models.

Reference	Model
Ali et al. [20]	$\tau_{\max} = 0.54\sqrt{f_c}b_p^{0.4}t_p^{0.3}$
Zhang et al. [21]	$\tau_{\max} = 1.15\gamma^{0.138}f_c^{0.613}$
Seracino et al. [16]	$\tau_{\max} = (0.802 + 0.078\varphi_f^{0.526})f_c^{0.6}$
Seracino et al. [19]	$P_{IC} = \sqrt{\tau_{\max}\delta_{\max}}\sqrt{L_{\text{per}}(EA)_P}$

A total of 246 test data were collected and divided into five sets in terms of their failure modes (as shown in Tables 2–6). As for the fiber type, 5 different fiber were involved, such as carbon FRP (CFRP), aramid FRP (AFRP), glass FRP (GFRP), basalt FRP (BFRP), and graphite FRP.

## 4. Model Uncertainty Factor

Model uncertainty inevitably exists due to varying degrees of idealization involved in calculation methods. The error between the predicted strength and the experimental value was described by a model uncertainty factor, which can be defined by the following formula [38, 39]:

$$P_u^m = \varepsilon \times P_u^c. \quad (1)$$

$P_u^m$  herein is the actual measuring bonded joints strength,  $P_u^c$  is the calculated prediction value, and  $\varepsilon$  is the model uncertainty factor which can quantitatively evaluate the impact of model uncertainty on the structure. A model uncertainty factor greater than 1 means that the measured value is larger than the calculated strength, and vice versa. Theoretically, the best value of  $\varepsilon$  is 1.0; therefore, the value getting closer to 1.0 suggests a more accurate model. The coefficient of variance (CV) of  $\varepsilon$  indicates the dispersion degree of prediction models. The smaller the CV of  $\varepsilon$  is, the higher the accuracy of the model is.

Practically,  $\varepsilon$  greater than 1 is regarded to be conservative, while the opposite is unsafe. On the basis of the collected pullout test data, the  $P_u^c$  under different failure modes could be calculated. Then, the model uncertainty factor  $\varepsilon$  could be obtained according to equation (1).

Figures 1–5 plot the contrast between the calculated value  $P_u^c$  (vertical axis) and the experimental value  $P_u^m$  (horizontal axis) for the three models under five failure modes. If the data points are tightly distributed near the 45-degree line, the model is considered to be accurate. Conversely, the model has a great dispersion. As shown in Figures 1–5, for the three models under the five failure modes, their data points are not distributed near the 45-degree line, indicating that these three models are very discrete.

Tables 7–11 display the statistics of calculated  $\varepsilon$  under five failure modes, including their mean value, the standard deviation (SD), and the coefficient of variation (CV). The tables below have shown the CV of  $\varepsilon$  ranging from 0.4 (the AM Model and Zhang Model under adhesive failure mode) to 0.61 (the three models under adhesive-concrete interface failure mode). However, the CV of a model uncertainty factor ranging from 0.2 to 0.3 is usually considered to be reasonable [40].

TABLE 2: Database under the adhesive-concrete interface failure mode.

Reference	Tests number	Material	$f_c$ (MPa)	$E_f$ (GPa)	$t_p$ (mm)	$t_f$ (mm)	$t_g$ (mm)
[13]	16	CFRP/GFRP	22	37.17–174.17	8–11.3	8–11.3	10–24
[24]	6	CFRP	34.5	124	12	12	20–25
[29]	13	BFRP/GFRP/CFRP	19	46–182	6–15	1.4–10	5–15
[30]	28	BFRP/GFRP/CFRP	19	42–182	6–15	2.5–10	8–20
[31]	11	BFRP/GFRP/CFRP	19.6–49.3	51–177	6–15	1.4–8	5–15
[32]	15	CFRP	34	124–210	12	12	20
Total	89	BFRP/GFRP/CFRP	19–49.3	37.17–210	6–15	1.4–12	5–25

TABLE 3: Database under the adhesive-FRP interface failure mode.

Reference	Number of tests	FRP material	$f_c$ (MPa)	$E_f$ (GPa)	$t_p$ (mm)	$t_f$ (mm)	$t_g$ (mm)
[33]	3	CFRP	36.199	37.17–174.17	11	11	14
[24]	2	CFRP	34.5	124	12	12	25
[29]	7	GFRP/CFRP	19	46–182	8	8	14–15
[30]	4	GFRP/CFRP	19	51–159	8–10	8–10	14–15
[34]	6	BFRP/GFRP/CFRP	36.2	40.8–155	6–10	6–10	5–15
[35]	6	CFRP	34.86	124–165	8–20	1.4–8	6–20
[31]	1	CFRP	19.6	155	8	8	14
[36]	2	CFRP	28.5	149–155	9.5	9.5	14.25–19
Total	31	BFRP/GFRP/CFRP	19–36.2	37.17–182	6–20	1.4–12	5–25

TABLE 4: Database under the adhesive failure mode.

Reference	Number of tests	FRP material	$f_c$ (MPa)	$E_f$ (GPa)	$t_p$ (mm)	$t_f$ (mm)	$t_g$ (mm)
[13]	2	CFRP	22	37.17	11	11	14–18
[31]	9	BFRP/GFRP/CFRP	19.6–49.3	46–182	6–15	2.5–8	8–14
[36]	16	GFRP/CFRP	28.5	41–155	7.5–16	4.5–10	8–20
[37]	13	GFRP/CFRP	41	42–139	9.5–12.7	9.5–12.7	14.25–25.4
[6]	2	CFRP	34–65.7	146	12	12	20
[16]	3	GFRP	33.4–49.2	161.8–162.3	14.9–20.37	1.26–1.28	3.26–3.27
Total	45	BFRP/GFRP/CFRP	19–65.7	37.17–182	6–20.37	1.26–12.7	3.26–25.4

TABLE 5: Database under the concrete crush or FRP rupture failure mode.

Reference	Number of tests	FRP material	$f_c$ (MPa)	$E_f$ (GPa)	$t_p$ (mm)	$t_f$ (mm)	$t_g$ (mm)
[33]	4	CFRP	34.47–36.199	227.54	11	11	14
[16]	10	GFRP	33.4–53	161.8–162.3	10.1–20.47	1.24–1.28	3.24–3.28
[29]	4	BFRP/CFRP	19	46–177	6–10	1.4–6	10
[30]	4	BFRP/CFRP	19	46–177	6–10	1.4–6	5–10
[37]	8	CFRP/GFRP	41	43.5–128	9.5–12.7	9.5–12.7	14.25–25.4
[6]	4	CFRP	34–65.7	146	12	12	30
[31]	4	CFRP/GFRP/BFRP	29.7	48–180	8–10	8–10	13–15
Total	38	CFRP/GFRP/BFRP	19–65.7	43.5–227.54	6–20.47	1.24–12.7	3.24–30

TABLE 6: Database under the concrete surface failure mode.

Reference	Number of tests	FRP material	$f_c$ (MPa)	$E_f$ (GPa)	$t_p$ (mm)	$t_f$ (mm)	$t_g$ (mm)
[16]	12	CFRP	38.8	140–170	10	3.6–6	20
[19]	9	GFRP	30–64.8	144.6–162.3	9.95–20.23	1.2–2.97	3.26–3.27
Total	33	CFRP/GFRP	30–64.8	140–170	9.95–20.23	1.2–6	3.26–20

Therefore, for the models with quite high CV, such as the three models under the adhesive-concrete interface failure mode with a CV of 0.61 as listed in Table 7, a

further investigation into the systematic reason which causes a large prediction deviation was required [41]. That is to say, the model uncertainty factor  $\varepsilon$  directly calculated

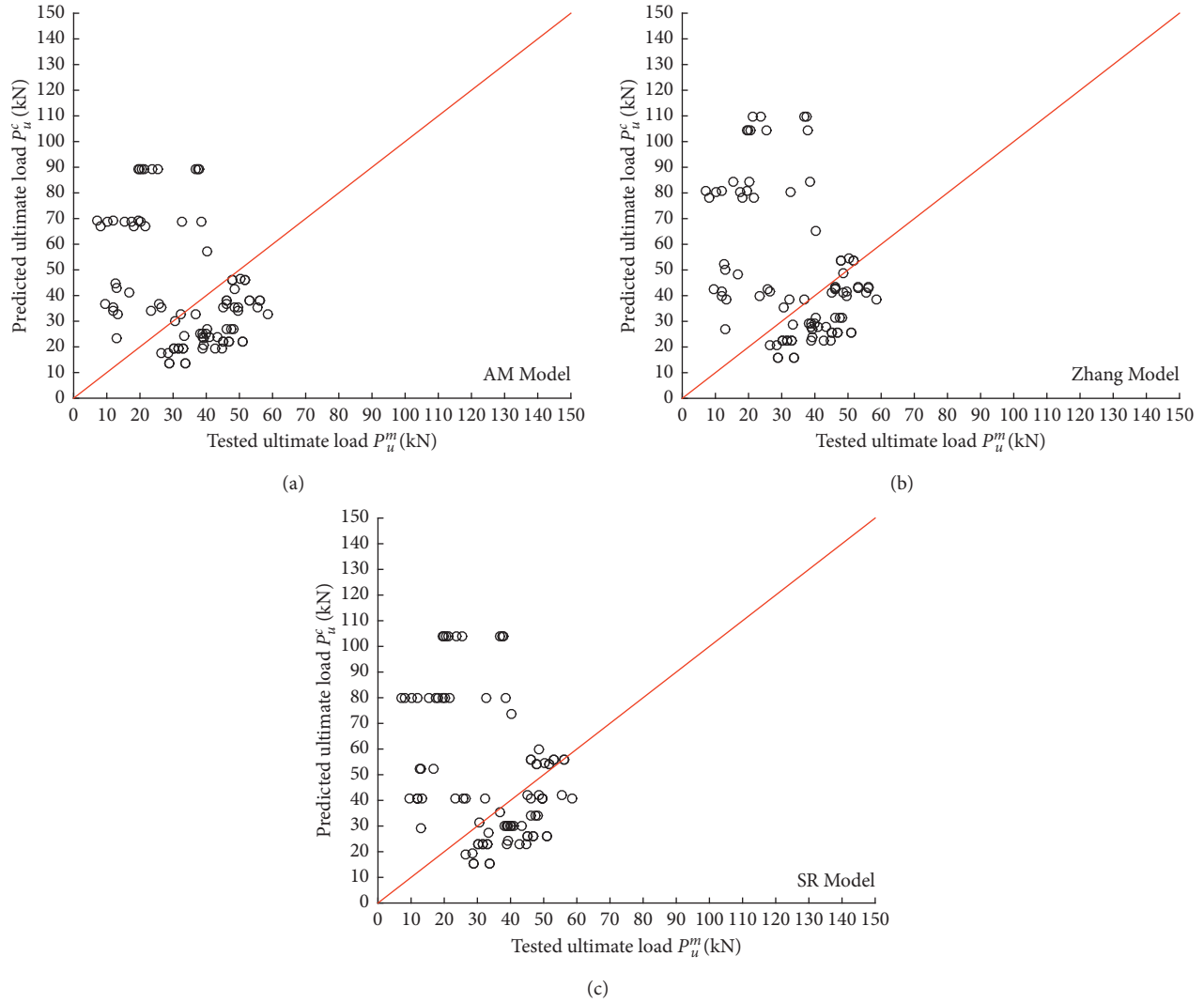


FIGURE 1: Comparison between the calculated value and the experimental value for the three models under the adhesive-concrete interface failure mode: (a) AM Model, (b) Zhang Model, and (c) SR Model.

by equation (1) was necessary to be checked for randomness.

Take the AM Model as an example, a scatter plot of the model uncertainty factor (written as  $\varepsilon_{AM}$ ) against the concrete compression strength  $f_c$  is shown in Figure 6, where an obvious nonlinear relationship existed, indicating that the model uncertainty factor  $\varepsilon$  calculated by equation (1) was not a random variable but with an obvious dependence upon the design parameters, i.e., the concrete compression strength  $f_c$  in this case.

In order to clarify whether the model uncertainty factor is systematically dependent on the design parameters, a correlation analysis is needed.

In this paper, the distribution of the model uncertainty factor is featured by uncertainty. Hence, the Spearman correlation coefficient method, which has no specific requirements on the distribution characteristics of the data, was adopted to perform a correlation analysis for the model uncertainty factor and the five input parameters.

When the significance level ( $p$  value) is larger than 0.05, the Spearman correlation coefficient method is a non-parametric test with a null hypothesis of zero-rank correlation. In a Spearman correlation analysis, a customary significance value is larger than 0.05, and the absolute value of  $r$  close to 1.0 means high dependence. The Spearman correlation analysis showed that there was a negative correlation between  $\varepsilon_{AM}$  and the design parameters, with a high level of  $r$ -value ranging from  $-0.450$  to  $-0.728$  and a low level of  $p$  value ranging from 0.001 to 0.013 (see the second and fourth columns in Table 12). It statistically demonstrated that the model uncertainty factor  $\varepsilon_{AM}$  was systematically dependent on the five input parameters.

Similarly, Spearman correlation analysis was performed for Zhang Model and SR Model, respectively. The results indicated that the uncertainty factors of all the three models were statistically dependent upon the design parameters. Therefore, a further analysis for the dependency was necessary.

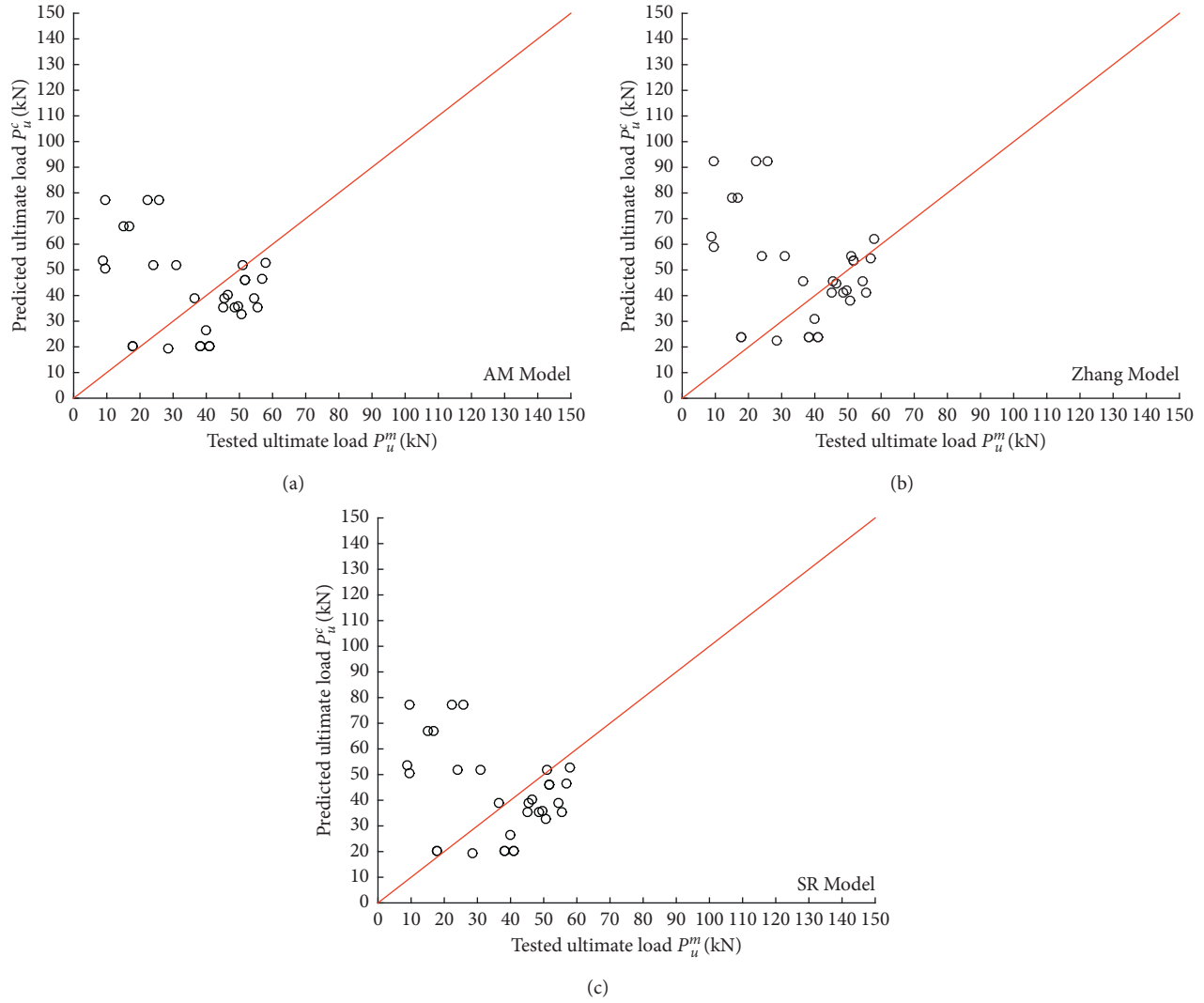


FIGURE 2: Comparison between the calculated value and the experimental value for the three models under the adhesive-FRP interface failure mode: (a) AM Model, (b) Zhang Model, and (c) SR Model.

By definition, since the model uncertainty factor must be a random variable with no dependency on the design parameters, a multiple regression analysis should be performed to eliminate the correlation between the model uncertainty factor and the design parameters, which is resulted from the systematic correlation part.

The residual random factor (represented by  $\varepsilon^*$ ) could be obtained by eliminating the systematic dependence (expressed by a multiple regression function  $f$ ). That is, the model uncertainty factor  $\varepsilon$  can be regarded as a composition of a systematic correlation part  $f$  and a residual random factor  $\varepsilon^*$ :

$$\varepsilon = f \times \varepsilon^*. \quad (2)$$

Substituting equation (2) into equation (1), equation (1) then turns into

$$P_u^m = \varepsilon^* \times f \times P_u^c. \quad (3)$$

Since the CV of the three selected models was high, the model uncertainty factor was then characterized by reducing the CV. Subsequently, a proper regression function  $f$  was

built and the residual random factor  $\varepsilon^*$  was characterized. Thus,  $\varepsilon^*$  was defined as the updated model uncertainty factor.

## 5. Model Modification

Regression analysis can explain the relationship between the model uncertainty factor and the design parameters by a regression equation which was established by collecting data points from test results. Then, the randomness of the model uncertainty factor and the precision of the regression equation were checked by the rest data points. In this study, all the three selected models under the five failure modes were calibrated, but only the AM Model under the adhesive-concrete interface failure mode was discussed in detail. In total, 89 data points have been collected for AM Model under adhesive-concrete interface failure mode (as shown in Table 2). 59 data points were used to establish the regression equation, and the rest 30 data points were used to check for randomness.

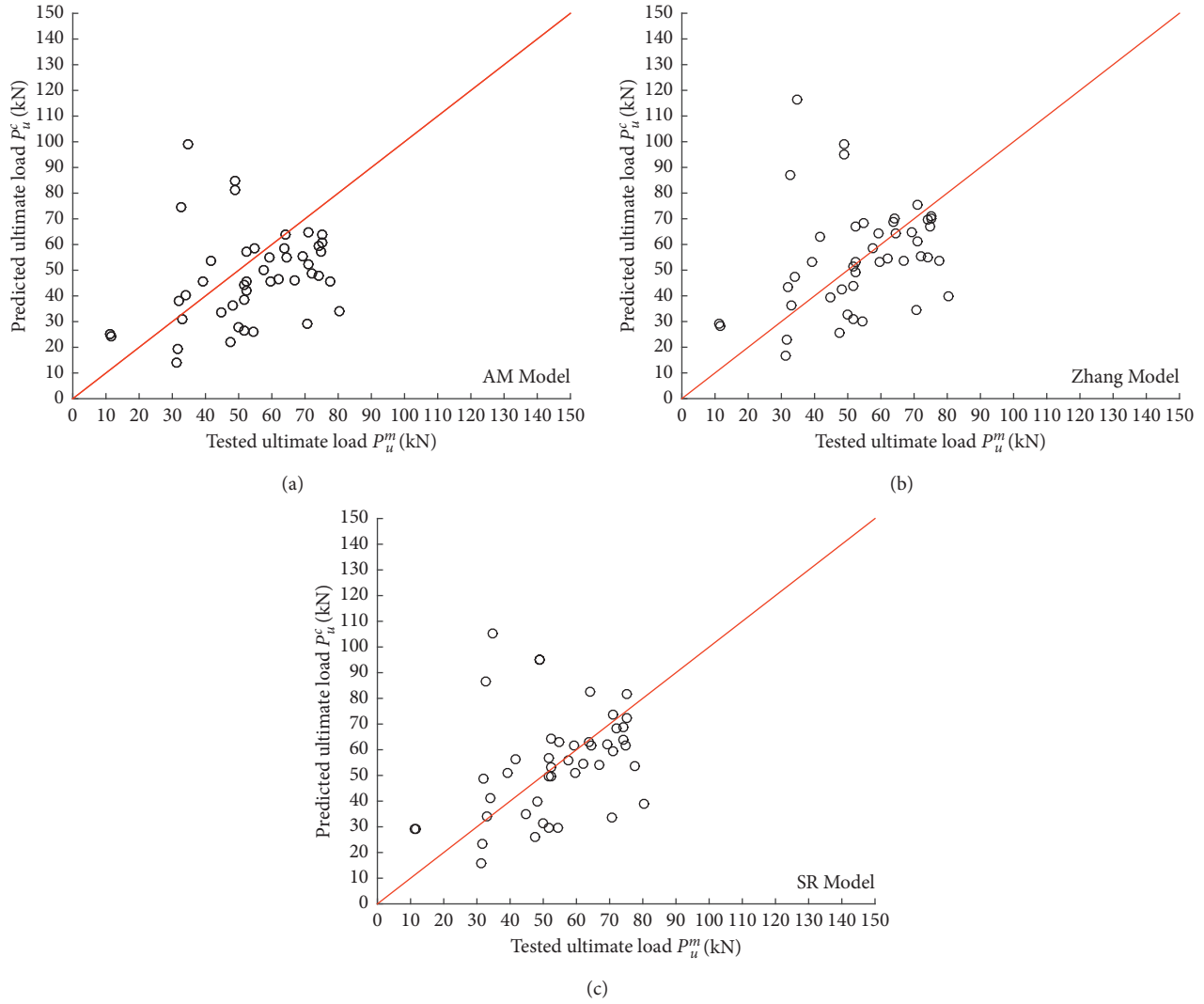


FIGURE 3: Comparison between the calculated value and the experimental value for the three models under the adhesive failure mode: (a) AM Model, (b) Zhang Model, and (c) SR Model.

The multiple regression analysis included two steps. In step one, the form of a correlation function (i.e., core function) was determined. The relationship between the model uncertainty factor and the design parameters can be expressed by the core function. The concerned functional relation equation was fitted by a MATLAB program written according to the damped LM algorithm [42, 43], and the calculated model uncertainty factors  $\varepsilon$  were plotted against the specific parameters by scatter plots in Figure 7.

The fitting function graphs show that the relation between the model uncertainty factor  $\varepsilon$  and each of the five design parameters was obviously nonlinear. The nonlinear trends with respect to the concrete compression strength  $f_c$  and the FRP height  $t_p$  were quite significant as shown in Figures 7(a) and 7(c). It can be noticed that the variation of  $\varepsilon$  with  $f_c$  and  $t_p$  can be more accurately fitted by a power function and an exponential function, respectively. For consistency reason, the exponential function and the power function were also adopted to fit the variation of  $\varepsilon$  with the

other three input parameters. Then, the core functions for the five input parameters can be gained as follows:

$$\varepsilon \propto f_c^{b_1}, \quad (4a)$$

$$\varepsilon \propto E_f^{b_2}, \quad (4b)$$

$$\varepsilon \propto e^{b_3 t_p}, \quad (4c)$$

$$\varepsilon \propto e^{b_4 t_f}, \quad (4d)$$

$$\varepsilon \propto e^{b_5 t_g}, \quad (4e)$$

where  $b_i$  is the regression coefficient. It must be noted that the regression coefficient  $b_i$  for each of the five core functions was still undetermined in the first step of regression. This is because the influence effect arising from the other four input parameters was represented by each  $b_i$  in the five equations above.

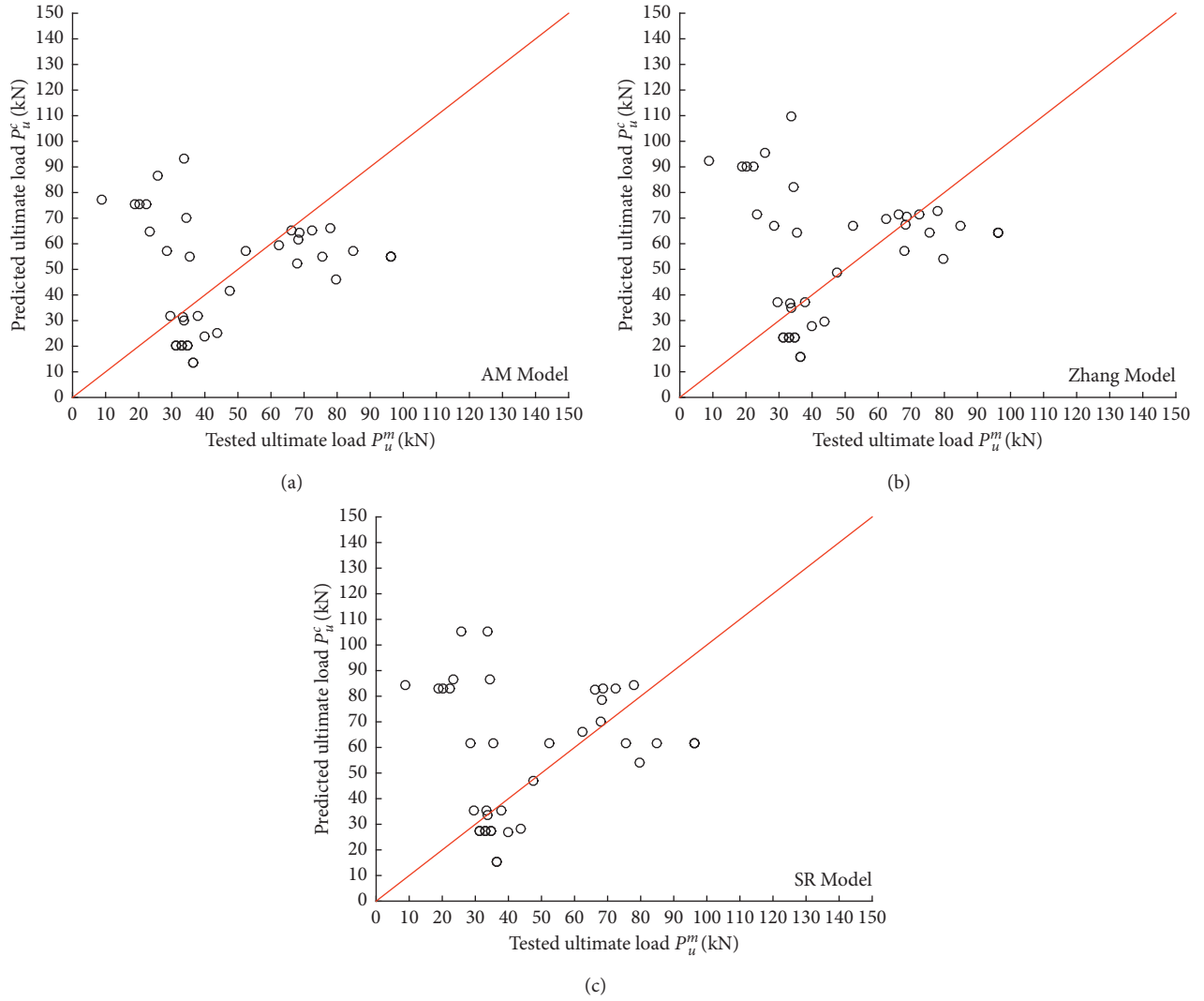


FIGURE 4: Comparison between the calculated value and the experimental value for the three models under concrete crush or FRP rupture failure mode: (a) AM Model, (b) Zhang Model, and (c) SR Model.

In the second step of regression, five core functions were combined together to constitute a multiplicative model  $f$ . Therefore, a regression function  $f$  was generated to multiplicatively describe the systematic dependence of the uncertainty factor on the five design parameters as follows:

$$f = e^{b_0} * e^{b_1 \ln f_c} * e^{b_2 \ln E_f} * e^{b_3 t_p} * e^{b_4 t_f} * e^{b_5 t_g}, \quad (5)$$

where  $b_i$  herein is the coefficient of the regression equation  $f$ . The model uncertainty factor can be given as follows:

$$\varepsilon = f * \varepsilon^* = e^{b_0} * e^{b_1 \ln f_c} * e^{b_2 \ln E_f} * e^{b_3 t_p} * e^{b_4 t_f} * e^{b_5 t_g} * \varepsilon^*. \quad (6)$$

$\varepsilon^*$  herein is the residual random factor, obtained by removing the correlation function  $f$  from the model uncertainty factor  $\varepsilon$ . The regression function can be transformed from a product form into a summation form through a logarithmic transformation on the two sides of equation (5). So the multiple nonlinear regression analysis

can be mathematically reduced to a multiple linear regression analysis.

The least square method was used to determine the five regression coefficients  $b_i$ . In our paper, a multiple linear regression analysis was carried out using SPSS to determine all the coefficients  $b_i$  as listed in Table 13. It can be seen from the table that each of the three models has a high determination of coefficient  $R^2$  (0.854 for the AM Model, 0.858 for the Zhang Model, and 0.840 for the SR Model).

According to the regression principles, the residual  $\varepsilon^*$  is a random variable with no dependence on the design parameters. However, it is necessary to check the residual  $\varepsilon^*$  for randomness by using a new set of test data. Hence, the remaining 30 data points were adopted here for the randomness verification.

With the AM Model as example, the results of Spearman correlation analysis are shown in Table 12, including the correlation coefficient  $r$  and the significance  $p$  value before and after the modification. The dependency of the model

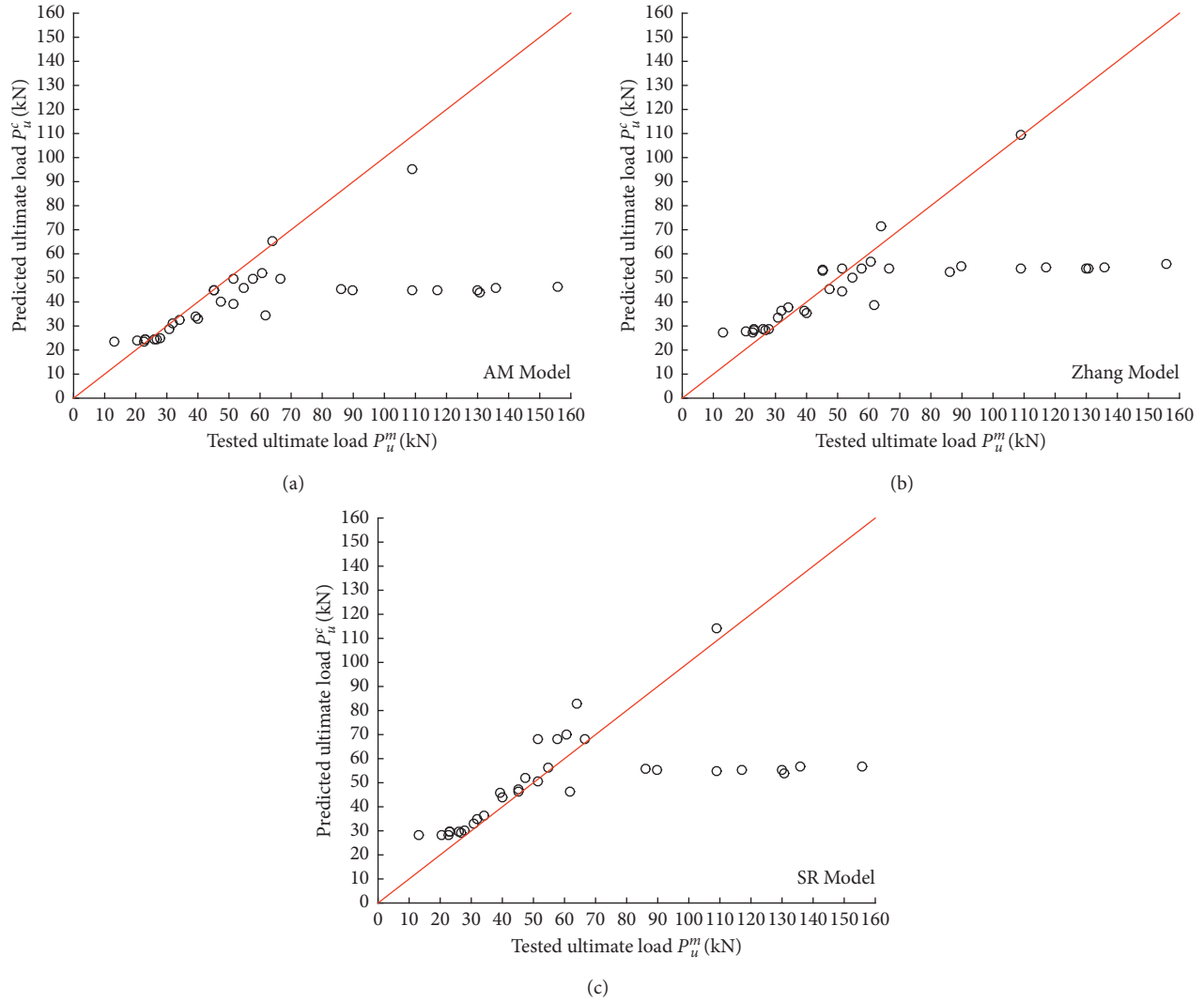


FIGURE 5: Comparison between the calculated value and the experimental value for the three models under concrete surface failure mode: (a) AM Model, (b) Zhang Model, and (c) SR Model.

TABLE 7: The statistics of calculated  $\varepsilon$  under adhesive-concrete interface failure mode.

Model	Mean	SD	CV
AM Model	1.17	0.7	0.61
Zhang Model	0.99	0.61	0.61
SR Model	0.96	0.59	0.61

uncertainty factor on the five input parameters was statistically proved to be sharply reduced. Thus, the residual factor  $\varepsilon^*$  can be regarded as a random part of  $\varepsilon$ .

Figure 8 plots the histogram of the residual part  $\varepsilon^*$  for each of the three modified models. As seen from the figure, for the three modified models, the mean value of  $\varepsilon^*$  was about 1.06, which was obviously closer to 1.00 in contrast to the original mean value of  $\varepsilon$ . Moreover, a mean value a little greater than 1.00 is regarded to be conservative and acceptable. Besides, for all the three models, the CV values of  $\varepsilon^*$  were markedly decreased to an acceptable 0.3, fully

TABLE 8: The statistics of calculated  $\varepsilon$  under adhesive-FRP interface failure mode.

Model	Mean	SD	CV
AM Model	1.05	0.58	0.55
Zhang Model	0.90	0.49	0.55
SR Model	1.05	0.58	0.55

demonstrating that the systematic correlation has been effectively eliminated by regression. Now, those three models' uncertainty factors have been adequately characterized.

The determined systematic correlation function  $f$  can be used for model modification, as shown below:

$$P_u^{c'} = f * P_u^c \quad (7a)$$

$P_u^{c'}$  herein is the predicted bonded joints strength after modification. The residual factor  $\varepsilon^*$ , a random variable, can



TABLE 9: The statistics of calculated  $\varepsilon$  under adhesive failure mode.

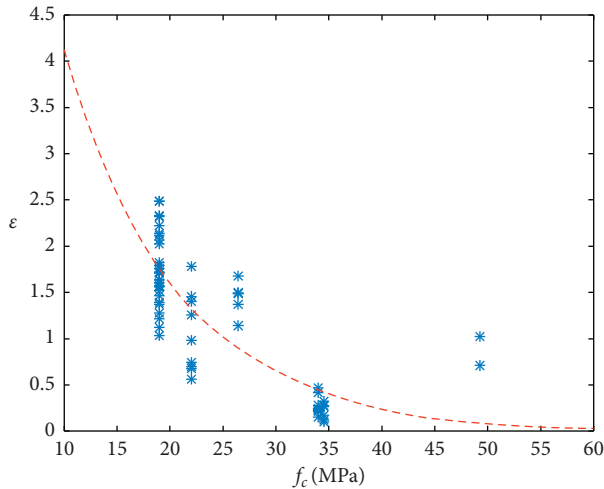
Model	Mean	SD	CV
AM Model	1.25	0.5	0.4
Zhang Model	1.08	0.43	0.4
SR Model	1.08	0.44	0.41

TABLE 10: The statistics of calculated  $\varepsilon$  under concrete crush or FRP rupture failure mode.

Model	Mean	SD	CV
AM Model	1.47	0.74	0.5
Zhang Model	1.26	0.6	0.48
SR Model	1.19	0.6	0.50

TABLE 11: The statistics of calculated  $\varepsilon$  under concrete surface failure mode.

Model	Mean	SD	CV
AM Model	1.17	0.62	0.53
Zhang Model	1.01	0.54	0.53
SR Model	0.98	0.54	0.55

FIGURE 6: The model uncertainty factor  $\varepsilon$  against concrete compression strength  $f_c$  (AM Model).

then be regarded as the new model uncertainty factor for modified prediction models:

$$P_u^m = \varepsilon^* * P_u^c. \quad (7b)$$

For all the three models, the comparison between the modified calculated value  $P_u^c$  and the measured value  $P_u^m$  was replotted (Figures 9(a)–9(c)). It is observed that all the modified data points are distributed near the 45-degree line after eliminating the systematic correlation.

In contrast with the original data (Figures 1(a)–1(c)), the difference between the calculated strength and the test value has been reduced significantly.

TABLE 12: Check for the randomness of the model uncertainty factor (AM Model).

Design parameter	Correlation coefficient $r$		Significance $p$ value	
	Before $\varepsilon$	After $\varepsilon^*$	Before $\varepsilon$	After $\varepsilon^*$
$f_c$	-0.728	0.090	0.001	0.634
$E_f$	-0.553	0.158	0.001	0.405
$t_p$	-0.450	0.318	0.013	0.087
$t_f$	-0.636	0.380	0.001	0.038
$t_q$	-0.528	0.31	0.003	0.096

## 6. Reliability Analysis

According to the formula put forward by the ACI code [44, 45], for the FRP NSM reinforcement system, the limit state function for a bonded joints strength design can be given as follows:

$$G = R - D - L. \quad (8)$$

$G$  herein is the limit state function,  $R$  is the bearing capacity,  $D$  is the dead load, and  $L$  is the live load. In a NSM FRP reinforced concrete structure, another factor, for example, the steel reinforcement, has also contributed much to the bearing capacity. However, as this part of resistance is hard to determine [30–32], only the FRP-concrete bonded joints resistance was therefore considered as the resistance in this paper [33], but not the contribution of steel reinforcement. Hence, for the NSM FRP reinforced concrete structure design, the capacity  $R$  is equivalent to the experimental value  $P_u$ :

$$R_d = P_u. \quad (9a)$$

The load combination can typically be presented as follows [46]:

$$S_d = 1.2D_n + 1.6L_n. \quad (9b)$$

$S_d$  herein is the design load, and  $D_n$  and  $L_n$  are the nominal dead load and live load, respectively. From the perspective of design, the nominal load is associated with the resistance. As a result, the nominal load can be represented by the resistance [36, 47]:

$$S_d = \varphi * P_u^c, \quad (9c)$$

where  $\varphi$  herein is the reduction factor for getting an appropriate reliability index [48, 49]. With no model uncertainty factor, the value of reduction factor will vary from 0.2 to 0.8 in different models [40]. However, by calibrating the model uncertainties, we could acquire a uniform value of reduction factor [34–37]. In this paper, for achieving an appropriate reliability index  $\beta$  (about 3.00), the value of the reduction factor  $\varphi$  was set uniformly as 0.6. Considering that the NSM FRP reinforcement system may be applied in different loading conditions, the live-to-dead load ratio  $\eta = L_n/D_n$  was set as 0.50, 0.75, 1.00, 1.25, and 1.50, respectively.

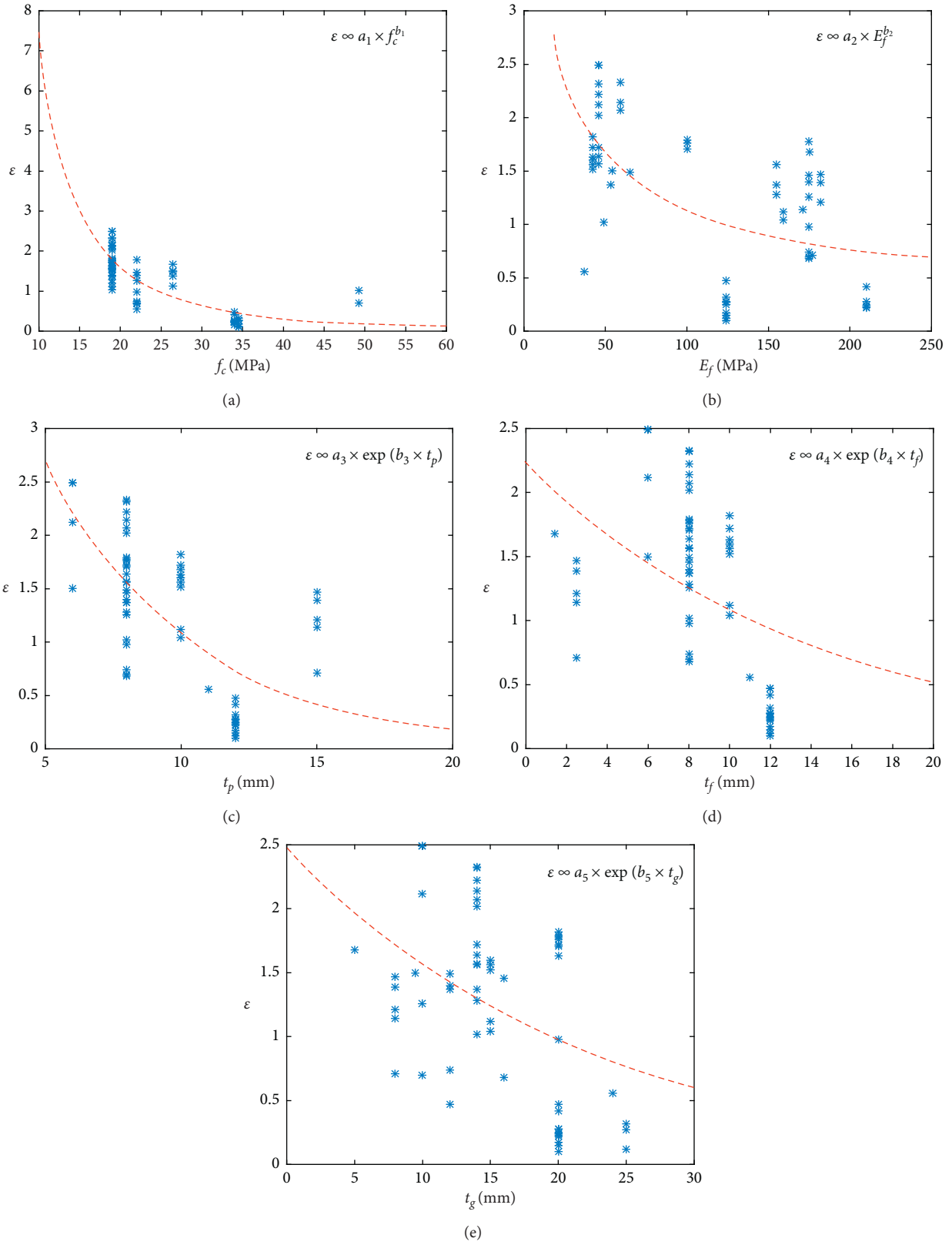


FIGURE 7: Fitting function graph: (a) concrete compression strength, (b) FRP modulus of elasticity, (c) FRP height, (d) FRP thickness, and (e) groove width.

TABLE 13: Coefficients in the regression function  $f$  for the three models.

Coefficients	AM Model	Zhang Model	SR Model
$b_0$	7.494	7.389	7.686
$b_1$	-1.245	-1.256	-1.036
$b_2$	-0.327	-0.331	-0.364
$b_3$	-0.090	-0.087	-0.114
$b_4$	-0.150	-0.154	-0.140
$b_5$	0.003	0.003	-0.005
$R^2$	0.854	0.858	0.840

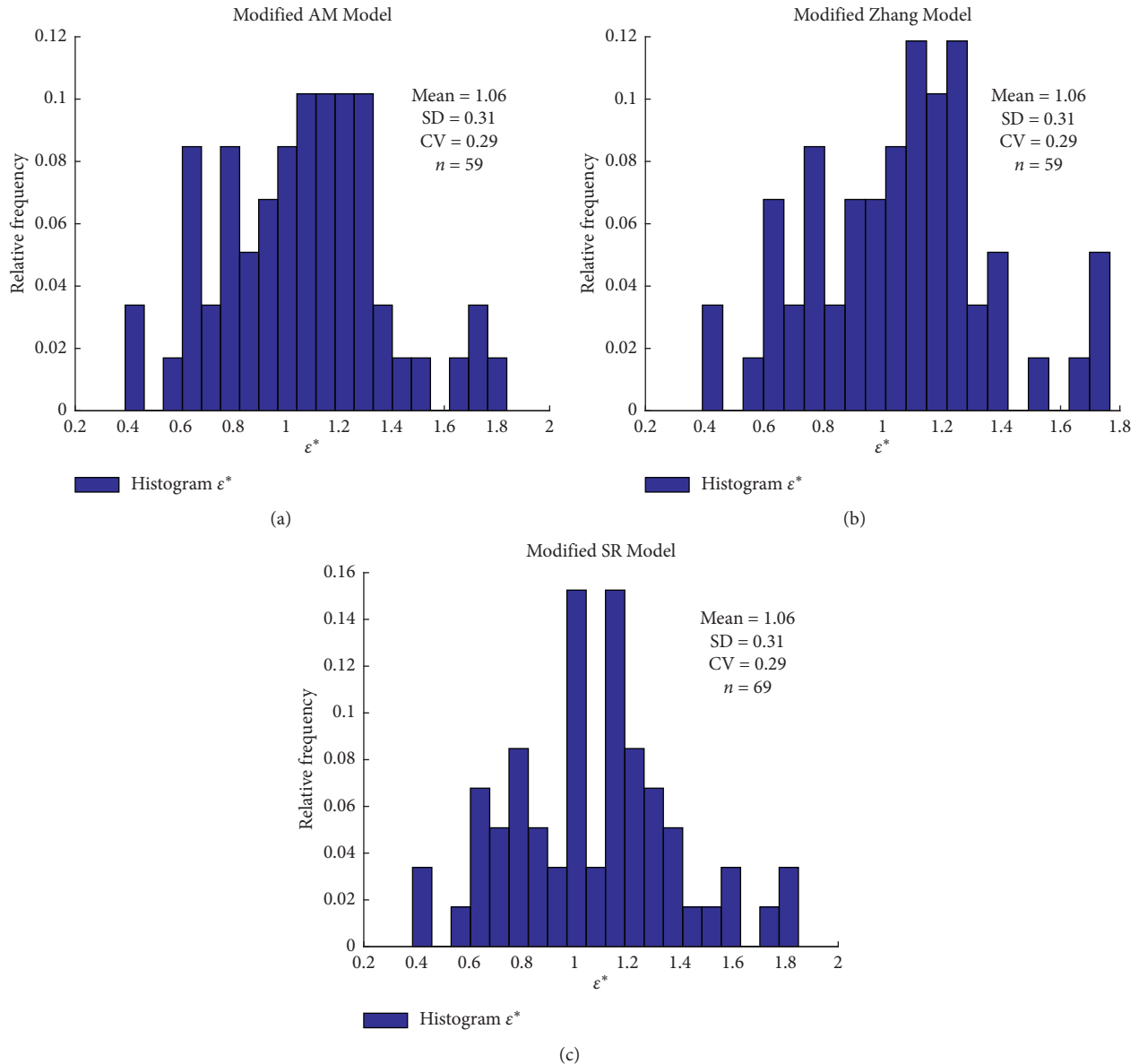


FIGURE 8: Histogram of  $\epsilon^*$  for each modified model: (a) AM Model, (b) Zhang Model, and (c) SR Model.

Obviously, the five design parameters (i.e.,  $f_c$ ,  $E_f$ ,  $t_p$ ,  $t_f$  and  $t_g$ ) were the main influence factors in the reliability analysis. In Table 14, two groups of commonly used nominal parameters (A and B) were included for reliability analysis [50].

These two groups of parameters were from the literature review and were frequently used in the NSM FRP-concrete bonded joints reliability analysis. A set of groove width  $t_g$  was selected from the references and experiments (i.e., 21.3,

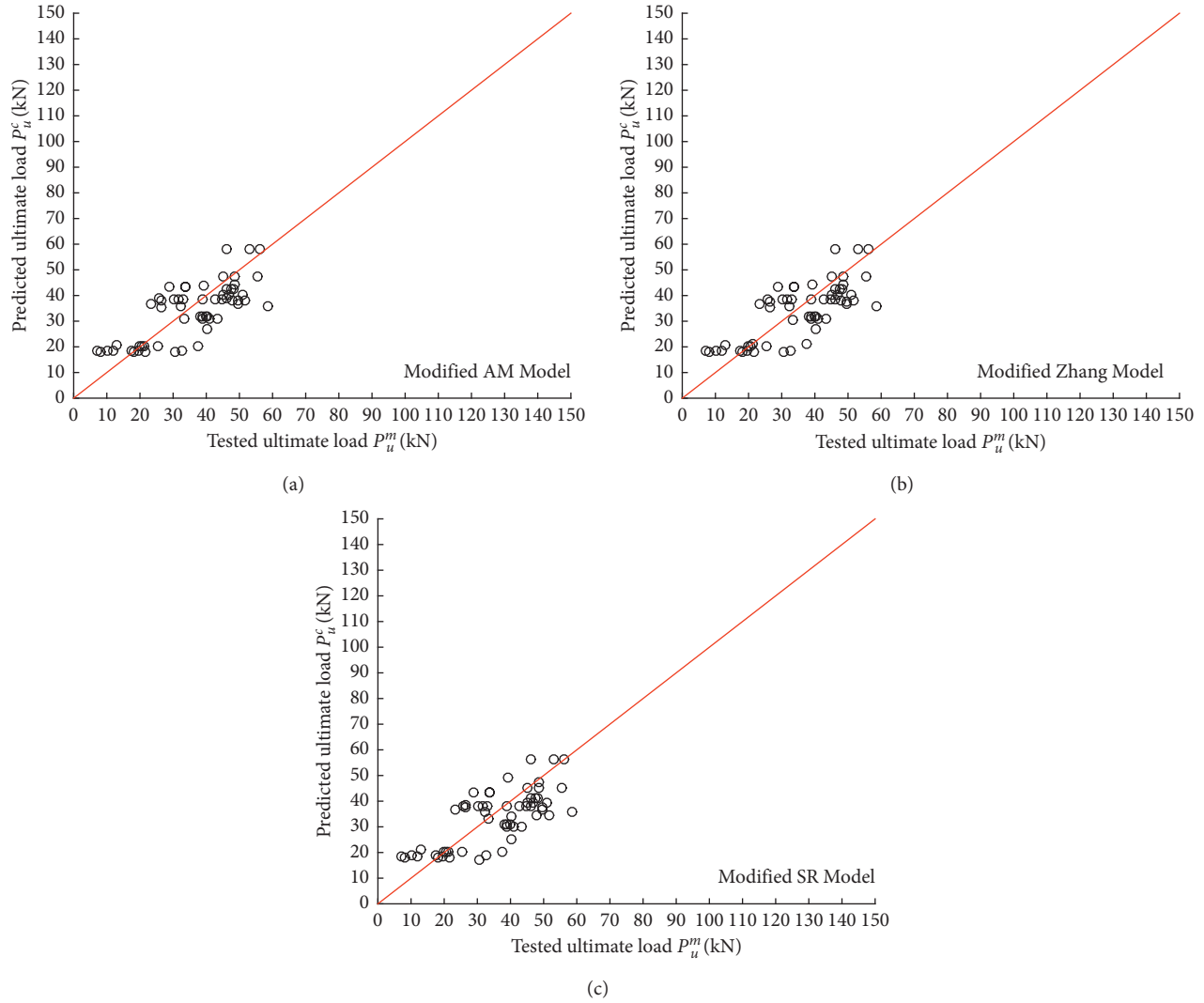


FIGURE 9: Comparison between the calculated value and the experimental value for the three modified models under adhesive-concrete interface failure mode: (a) AM Model, (b) Zhang Model, and (c) SR Model.

TABLE 14: Two groups of nominal parameters for the reliability analysis.

Design parameter	Nominal value A	Nominal value B	Reference
$f_c$ (MPa)	27.56	41.3	[35, 36]
$E_f$ (GPa)	52	165	[37]
$t_p$ (mm)	4	8	[35]
$t_f$ (mm)	8	12	[50], measured date
$t_g$ (mm)	19	9	Measured date
$\bar{D}$ (kN)	Equation (9)	Equation (9)	Galambos et al. [51]
$L$ (kN)	Equation (9)	Equation (9)	Galambos et al. [51]

22.33, 21.87, 23.19, 20.53, 21.57, 19.53, 20.30, 19.77, 19.67, 19.80, and 22.97). The nominal  $t_g$  was set to 19 based on the 95% probability of these values. Each of the other four design parameters (i.e.,  $f_c$ ,  $E_f$ ,  $t_p$ , and  $t_f$ ) was selected from one of the two groups (A and B). Thus, a sample space was generated by the five design parameters, with a sample size of  $1 \times 2 \times 2 \times 2 \times 2 = 16$ . Taking each of the five  $\eta$  (i.e., 0.50, 0.75,

1.00, 1.25, and 1.50) into account, we had  $16 \times 5 = 80$  cases of reliability analysis. For every case, JC method was applied for the calculation of the reliability index  $\beta$  [51–53].

Table 15 lists out three groups of reliability index  $\beta$  calculated in 16 design cases. In group 1, model uncertainty factor was not taken into account. In group 2, model uncertainty factor was considered, but its systematic

TABLE 15: Reliability index of 16 cases for AM Model.

Case no.	Group 1: without considering the model uncertainty factor	Group 2: with the unmodified model uncertainty factor considered	Group 3: with the modified model uncertainty factor considered
1	2.58	1.53	3.15
2	3.56	1.62	3.05
3	4.82	2.36	2.98
4	3.43	2.47	3.12
5	0.78	1.64	3.48
6	1.39	1.35	3.23
7	3.47	2.94	3.78
8	1.09	3.08	3.09
9	6.62	1.51	2.95
10	4.22	1.73	2.91
11	2.45	1.61	3.96
12	4.58	1.93	4.03
13	5.12	1.54	3.26
14	3.27	1.26	3.01
15	4.40	3.13	2.87
16	5.03	2.98	2.95
Average	3.55	2.04	3.27
CV of $\beta$	0.44	0.32	0.11

dependence being not eliminated (abbreviated as unmodified model). In group 3, the model uncertainty factor was considered, as well as its systematic dependence being eliminated (abbreviated as modified model).

Take the AM model as an example, with the live-to-dead load ratio  $\eta$  at 1.00, the calculated reliability index  $\beta$  in 16 design cases is listed in Table 15 [6, 38]. In group 1 (no model uncertainty factor), the calculated reliability index  $\beta$  had an average value of 3.5, which was the highest among the three groups. In group 2 (unmodified model) and group 3 (modified model), the average calculated reliability index  $\beta$  decreased to 2.04 and 3.27, respectively. In group 3 (modified model), the CV of  $\beta$  has a lowest value of 0.11, which was clearly less than the other two groups (0.32 and 0.44).

Generally, when the uncertainty is significantly underestimated, the reliability of design considering no model uncertainty factor is very unsafe [29, 43]. Our result also indicated that very different reliability did exist in the same design case, which is due to the systematic correlation of  $\varepsilon$ . While practically, the same input parameters leading to quite different output reliability levels are unreasonable [42]. In summary, by eliminating the systematic correlation of  $\varepsilon$ , all the problems above can get effective solutions.

## 7. Conclusions

Some prediction models have been proposed for NSM FRP-concrete bonded joints strength. Model uncertainty inevitably does exist due to limited experimental data, incomplete research parameters, and idealized calculation method. A method of model calibration for these prediction models is presented in this study. The main conclusions were summarized as follows:

- (1) A total of 246 pullout test data were collected to calibrate three selected prediction models. The model uncertainty factor was defined to quantitatively evaluate the uncertainties in a model. By using

the Spearman correlation analysis, the model uncertainty factor calculated by a prediction model was checked for randomness and was proved to have a strong dependency on the design parameters.

- (2) A multiple regression analysis was applied and a regression equation was established to reduce the value of CV for the model uncertainty factor. The systematical dependence of the model uncertainty factor on design parameters was then eliminated, and the residual factor after regression was checked for randomness. After modification, the model uncertainty factors of the three selected models have become reasonable random variables which followed the logarithmic normal distribution.
- (3) For different NSM FRP bonded joints strength models, the model uncertainties can be brought to the same level after model modification. The model uncertainty factor after modification was appreciated for its merit by performing a reliability analysis using the JC method, and the calibration significantly increased the accuracy of the prediction models.

This study has widened and deepened the knowledge of the NSM FRP interfacial bonded joints strength prediction models and is desirable for guidelines to standardize the calibration of model uncertainties.

## Data Availability

Data in this article used to support this study are currently under embargo while the research findings are commercialized. Requests for data will be considered 6 months after the publication of this article by the corresponding author.

## Conflicts of Interest

The authors declare that they have no conflicts of interest regarding this work.

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